

# STRONG LENS TIME DELAY ESTIMATION

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5. Bayesian: Full posterior and sampler
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  - ▶ Problem
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11. (If time allows) A new MCMC method for multimodality

# INTRODUCTION

## Strong gravitational lensing

Credit: NASA's Goddard Space Flight Center

The strong gravitational field of a lensing galaxy splits light into two images.

- ▶ Light rays take **different routes** whose **lengths can be different**.
- ▶ **Difference between their arrival times** → **Time delay ( $\Delta$ )**

# INTRODUCTION

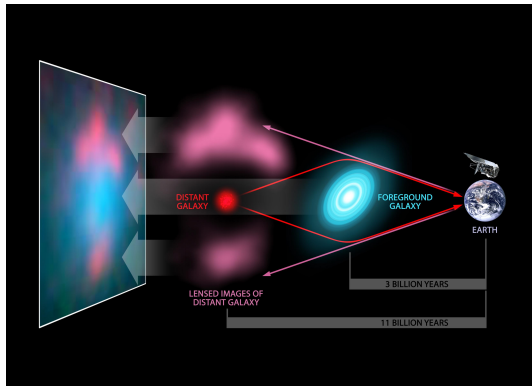



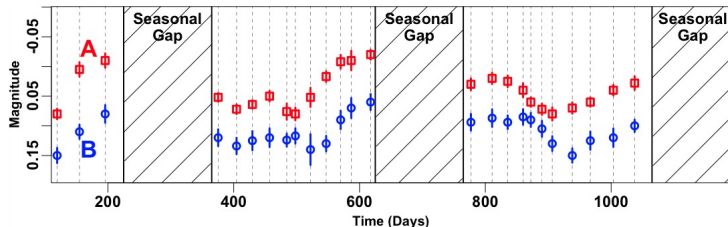
Image Credit: NASA/JPL-Caltech

Time delay is used to infer cosmological parameters, e.g.,

- ▶ Hubble constant  $H_0$  (Refsdal, 1964)
- ▶ Equation of state of dark energy (Linder, 2011)

# DATA

Simulated data of a **doubly-lensed** quasar , an active black hole (Image Credit: NASA/ESA).

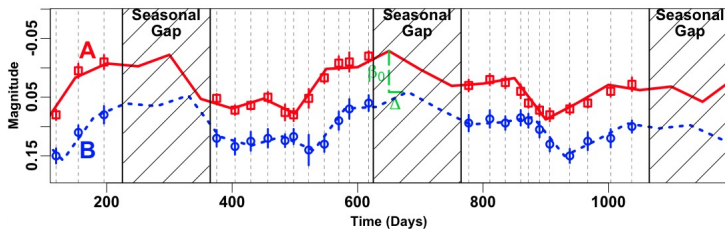


Data are composed of two time series with measurement errors.

- ▶ Observation times  $\mathbf{t} \equiv \{t_1, t_2, \dots, t_n\}^\top$
- ▶ Observed magnitudes  $\mathbf{x} \equiv \{x_1, x_2, \dots, x_n\}^\top$ , and  $\mathbf{y}$
- ▶ Measurement errors (SD)  $\boldsymbol{\delta} \equiv \{\delta_1, \delta_2, \dots, \delta_n\}^\top$  and  $\boldsymbol{\eta}$

Our job is to estimate **time delay** (shift in the horizontal axis) between two time series.

# STATE-SPACE MODEL



- ▶  $\exists$  latent light curves representing the unobserved true magnitudes in continuous time (red and blue dashed curves).

$\mathbf{X}(t) = (X(t_1), X(t_2), \dots, X(t_n))^T$  and  $\mathbf{Y}(t)$ , values on curves at  $t$

- ▶ A curve-shifted model (Pelt et al., 1994):

$$\mathbf{Y}(t) = \mathbf{X}(t - \Delta) + \beta_0,$$

where the time delay  $\Delta$  and magnitude offset  $\beta_0$  are unknown.

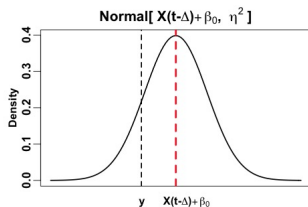
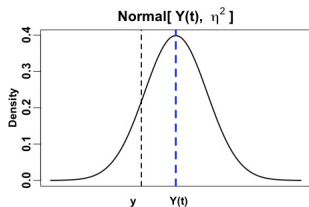
# DISTRIBUTIONS OF THE OBSERVED DATA

Observed data with independent Gaussian measurement errors

▶  $x_j \mid \mathbf{X}(t_j) \stackrel{\text{indep.}}{\sim} \text{Normal}[\mathbf{X}(t_j), \delta_j^2]$

▶  $y_j \mid \mathbf{Y}(t_j) \stackrel{\text{indep.}}{\sim} \text{Normal}[\mathbf{Y}(t_j), \eta_j^2]$

$y_j \mid \mathbf{X}(t_j - \Delta), \Delta, \beta_0 \stackrel{\text{indep.}}{\sim} \text{Normal}[\mathbf{X}(t_j - \Delta) + \beta_0, \eta_j^2].$



▶  $p(\mathbf{x}, \mathbf{y} \mid \mathbf{X}(\mathbf{t}^\Delta), \Delta, \beta_0)$

$$= \prod_{j=1}^n p[x_j \mid \mathbf{X}(t_j)] \times p[y_j \mid \mathbf{X}(t_j - \Delta), \Delta, \beta_0],$$

where  $\mathbf{t}^\Delta$  denotes the sorted  $2n$  observation times of  $\mathbf{t}$  and  $\mathbf{t} - \Delta$ .

# DISTRIBUTIONS OF THE LATENT DATA

Latent data with Ornstein-Uhlenbeck (O-U)/Damped RW process

- ▶ Many astrophysicists have supported the O-U process; Kelly+ (2009), Kozłowski+ (2010), MacLeod+ (2010), Zu+ (2013), Tewes+(2013), Hojjati+(2014), Bonvin+(2016), and more!
- ▶  $d\mathbf{X}(t) = -\frac{1}{\tau}(\mathbf{X}(t) - \boldsymbol{\mu})dt + \sigma dB(t)$ , where  $\tau$  is a mean-reversion time,  $\boldsymbol{\mu}$  is the overall mean, and  $\sigma$  is the short-term variability.
- ▶ O-U process is a Gaussian process with a Matérn(1/2) kernel.
- ▶  $\mathbf{X}(t^\Delta) \sim \text{Normal}_{2n}$  with some mean vector and covariance matrix
- ▶  $p(\mathbf{X}(t^\Delta) \mid \boldsymbol{\mu}, \sigma, \tau, \Delta) = p(\mathbf{X}(t_1^\Delta) \mid \boldsymbol{\mu}, \sigma, \tau, \Delta) \times \prod_{j=2}^{2n} p(\mathbf{X}(t_j^\Delta) \mid \mathbf{X}(t_{j-1}^\Delta), \boldsymbol{\mu}, \sigma, \tau, \Delta)$



# BAYESIAN: PRIOR DISTRIBUTIONS FOR PARAMETERS

From statistician's perspective,

- ▶ Prior distributions must guarantee posterior propriety, i.e.

$$\int p(\text{parameters} \mid \text{data}) d(\text{parameters}) \\ = \int \text{Lik}(\text{parameters}) \times p(\text{parameters}) d(\text{parameters}) < \infty.$$

- ▶ Without knowing posterior propriety, no one can tell whether the resulting posterior sample is from the target posterior distribution or not (Hobert and Casella, 1994).
- ▶ When we are not sure about posterior propriety: Use proper priors!

From astrophysicist's perspective,

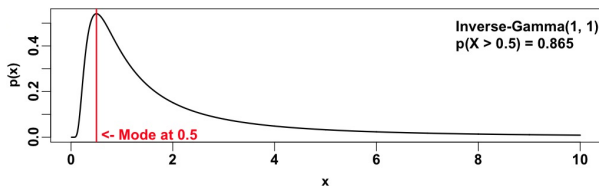
- ▶ Set up parameters in the proper prior distributions in a way to reflect on astrophysics and the dynamic of the O-U process.

# BAYESIAN: PRIOR DISTRIBUTIONS FOR PARAMETERS

- ▶  $\Delta \sim \text{Uniform}(u_1, u_2)$ 
  - ▶  $[u_1, u_2]$ , if  $\exists$  prior information to restrict the range of  $\Delta$ , e.g., a physical model of the lens, redshift, and relative locations.
  - ▶  $[t_1 - t_n, t_n - t_1]$ , otherwise.
- ▶ Mean of the O-U,  $\mu \sim \text{Uniform}(-30, 30)$ , why 30?
- ▶ Magnitude offset,  $\beta_0 \sim \text{Uniform}(-60, 60)$ , why 60?

Inv-Gamma distribution sets a **soft lower bound of a random variable**

If  $X \sim \text{Inv-Gamma}(a, b)$ ,  $p(x) \propto \frac{1}{x^{a+1}} \exp(-b/x)$  with a mode at  $\frac{b}{a+1}$ .



- ▶ Variance of the O-U,  $\sigma^2 \sim \text{Inv-Gamma}(1, c)$ , why  $1 \& c$ ?
- ▶ Timescale of the O-U,  $\tau \sim \text{Inv-Gamma}(1, 1)$ , why  $1 \& 1$ ?
- ▶ Estimates for 9,275 SDSS quasar light curves (MacLeod+, 2010)

# BAYESIAN: FULL POSTERIOR AND SAMPLER

Notation:  $\theta_{OU} \equiv (\mu, \sigma^2, \tau)$  and  $D_{obs} \equiv \{\mathbf{x}, \mathbf{y}\}$

- ▶ Full Posterior:  $p(\mathbf{X}(\mathbf{t}^\Delta), \Delta, \beta_0, \theta_{OU} \mid D_{obs})$ 
  - $\propto p(\mathbf{x}(\mathbf{t}), \mathbf{y}(\mathbf{t}) \mid \mathbf{X}(\mathbf{t}^\Delta), \Delta, \beta_0)$  Observed data
  - $\times p(\mathbf{X}(\mathbf{t}^\Delta) \mid \Delta, \theta_{OU})$  Latent data
  - $\times p(\Delta, \beta_0, \theta_{OU})$  Priors
- ▶ Metropolis-Hastings within Gibbs sampler
  1.  $p(\mathbf{X}(\mathbf{t}^\Delta), \Delta \mid \beta_0, \theta_{OU}, D_{obs}, )$
  2.  $p(\beta_0 \mid \mathbf{X}(\mathbf{t}^\Delta), \Delta, \theta_{OU}, D_{obs})$
  3.  $p(\theta_{OU} \mid \beta_0, \mathbf{X}(\mathbf{t}^\Delta), \Delta, D_{obs})$
- ▶ Pros: Complete investigation on all the model parameters
- ▶ Cons: Inefficient when  $\exists$  multimodality

# FREQUENTIST: PROFILE LIKELIHOOD

A profile likelihood function enables us to **focus on the parameter of interest** with nuisance parameters maximized out (Berger et al., 1996)

- ▶  $L(\Delta, \beta_0, \theta_{OU}) = \int_{\mathbf{R}^{2n}} p(\mathbf{x}, \mathbf{y} \mid \mathbf{X}(\mathbf{t}^\Delta), \Delta, \beta_0) \times p(\mathbf{X}(\mathbf{t}^\Delta) \mid \Delta, \theta_{OU}) d\mathbf{X}(\mathbf{t}^\Delta)$
- ▶  $L_{prof}(\Delta) \equiv \max_{\beta_0, \theta_{OU}} L(\Delta, \beta_0, \theta_{OU}) = L(\Delta, \hat{\beta}_0(\Delta), \hat{\theta}_{OU}(\Delta))$
- ▶  $L_{prof}(\Delta) \propto p(\Delta \mid D_{obs})$  asymptotically
- ▶ Provides approximate posterior mean, mode, standard deviation, and most importantly **shape of the (approximate) distribution**
- ▶ **Pros:** Simple to implement and easy to find multi-modes
- ▶ **Cons:** Computationally expensive for drawing a finer curve / no information about the relationship among parameters

# OUR TIME DELAY ESTIMATION STRATEGY

Bayesian and frequentist methods complement each other!

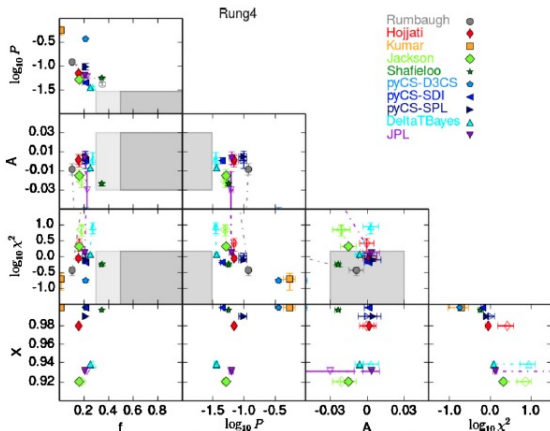
Our time delay estimation strategy:

1. Obtain the profile likelihood curve to check multimodality
2. Initialize Bayesian method near the modes identified by  $L_{\text{prof}}(\Delta)$

# EXAMPLE: TIME DELAY CHALLENGE

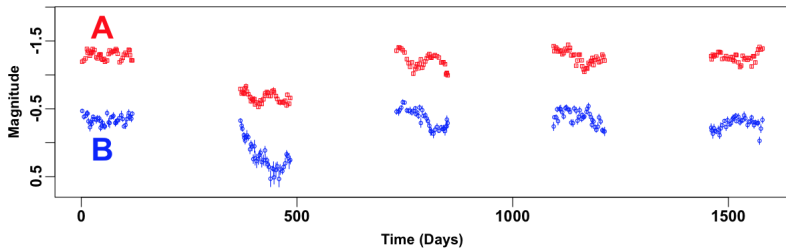
Time Delay Challenge (TDC, Dobler et al., 2015; Liao et al., 2015)

- ▶ A blind competition held by 8 astrophysicists from 2013 to 2014.
- ▶ Goals: (1) Providing an observation strategy for the LSST.  
(2) Improving current estimation methods.
- ▶ About 5,000 simulated data sets with some time delays (O-U).
- ▶ 13 teams blindly analyzed the simulated data.

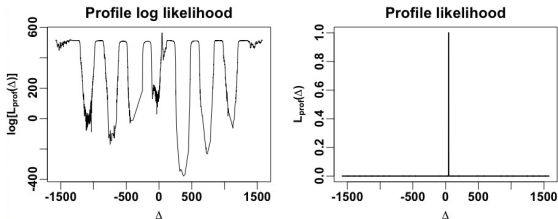


# EXAMPLE: TIME DELAY CHALLENGE

Simulated data of a doubly-lensed quasar from TDC

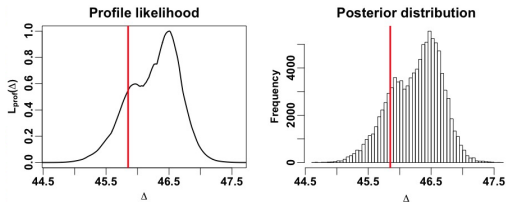


(1) The entire profile (log) likelihood curve



# EXAMPLE: TIME DELAY CHALLENGE

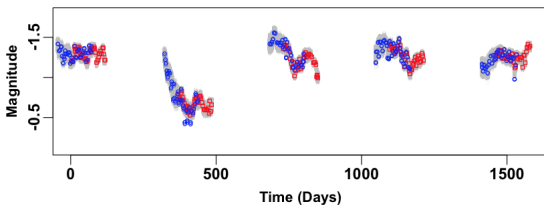
(2) Posterior distribution of  $\Delta$  initialized near the dominant mode



(3) Estimation summary for  $\Delta$

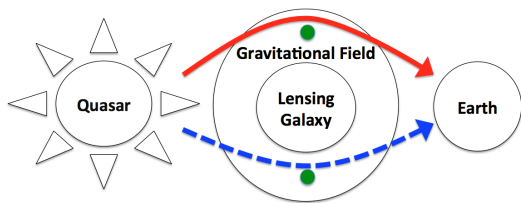
Method	Truth	Post. Mean	Post. SD
Bayesian	45.85	46.26	0.41
Profile likelihood		46.26	0.40

(4) Model Checking: Blue light curve is shifted by  $\hat{\Delta}$ , and  $\hat{\beta}_0$ .

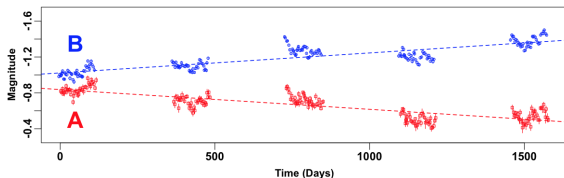




# MICROLENSING

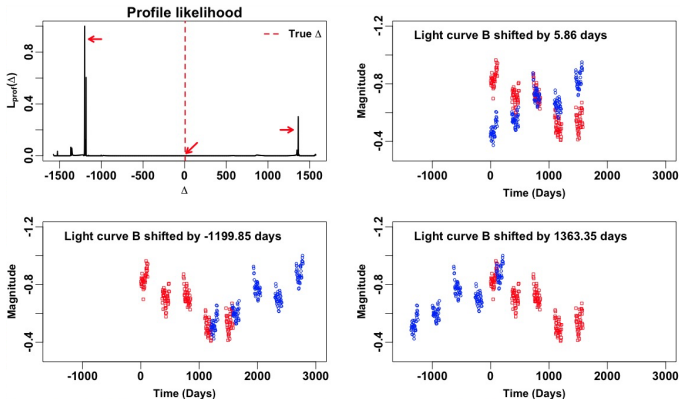


- ▶ **Microlensing** occurs when stars unusually close to the paths of light introduce **independent noise into magnification of brightness light curves** (Tewes et al., 2013).
- ▶ Two light curves may have **different long-term trends**, e.g., polynomial.



# MICROLENSING: PROBLEM

- ▶ A curve-shifted model does not work because one of the latent curves is no longer a shifted version of the other.



- ▶ A small overlap between two light curves (bottom plots) is the only similar fluctuation patterns detectable by shifting one of the light curves → several modes near margins of the range of  $\Delta$ .

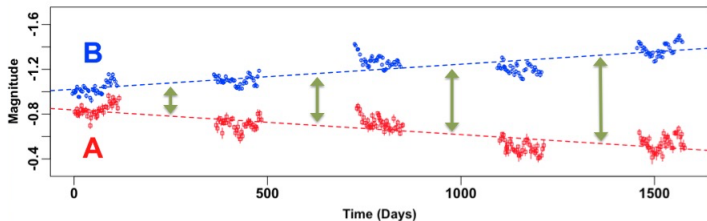
# MICROLENSING: MODEL

## Micro lensing model

- ▶ Popular way is to model the long-term trend of each light curve via a polynomial regression.
- ▶ Our microlensing model accounts for the difference between long-term trends using an  $m^{\text{th}}$ -order polynomial regression.

$$Y(t) = X(t - \Delta) + \mathbf{w}_m^T(t - \Delta)\beta,$$

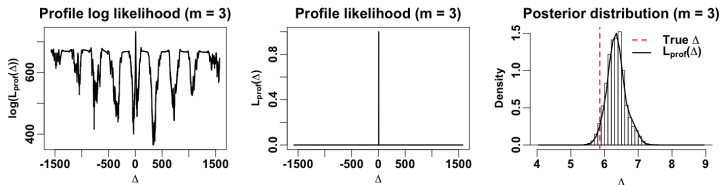
where  $\mathbf{w}_m(t - \Delta) \equiv (1, t - \Delta, \dots, (t - \Delta)^m)^\top$ , and  $\beta \equiv (\beta_0, \beta_1, \dots, \beta_m)^\top$  are regression coefficients.



- ▶ Our microlensing model reduces the number of regression coefficients by half!

# MICROLENSING EXAMPLE 1

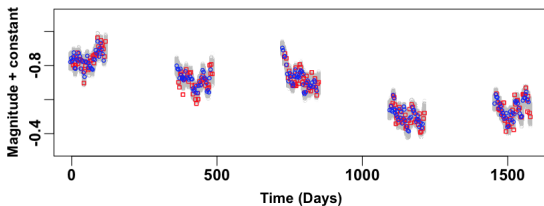
(1) We set  $m = 3$  as a default (Kochanek+, Morgan+, Tewes+).



(2) Estimation summary (Error  $\equiv |\Delta_{\text{true}} - \hat{\Delta}|$  and  $\chi \equiv \text{Error}/\text{SD}$ )

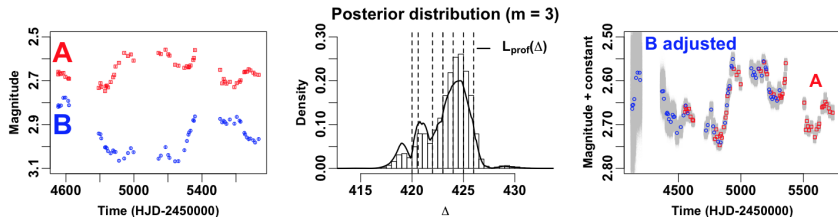
Method	Truth	E( $\Delta$  Data)	SD( $\Delta$  Data)	Error	$\chi$
Bayesian	5.86	6.34	0.28	0.48	1.71
Profile Lik.		6.36	0.28	0.50	1.76

(3) Model Checking: Blue light curve is adjusted by  $\hat{\Delta}$  and  $\hat{\beta}$ .



# MICROLENSING EXAMPLE 2: Q0957+561

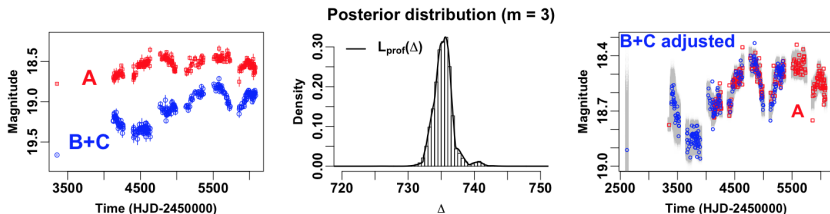
*r*-band data collected at the US Naval Observatory (Hainline+, 2012).



Researchers	Number of observations	Observation period	Measurement error (mag)	$\hat{\Delta}$	SE
Pelt et al. (1996)	831	1979–1994	0.0159	423	6
Oscz et al. (1997)	86	1994–1996	0.01, 0.02	424	3
Serra-Ricart et al. (1999)	197	1996–1998	0.023, 0.025	425	4
Oscz et al. (2001)	100	1994–1996	0.009, 0.01	423	2
Shalyapin et al. (2012)	371	2005–2010	0.012	420.6	1.9
This work	57	2008–2011	0.004	423.69 423.21	2.02 2.81

# MICROLENSING EXAMPLE 3: J1029+2623

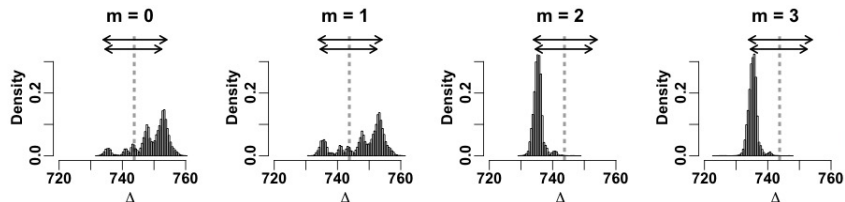
Data reported by Fohlmeister et al. (2013).



Researchers	Method	Estimate	90% Interval
Fohlmeister et al. (2013)	$\chi^2$ -minimization (AIC, BIC)	744	(734, 754)
Kumar, Stalin, and Prabhu (2014)	Difference-smoothing	743.5	(734.6, 752.4)
This work	Bayesian	735.28	(733.08, 737.59)
	Profile likelihood	733.11	(732.94, 738.44)

Our point estimate is much smaller than theirs (by about 10 days) & our 90% intervals are much shorter than theirs. What's wrong here? Is it because our model is over-confident?

## MICROLENSING EXAMPLE 3: J1029+2623 (CONT.)



Fohlmeister et al. (2013)

- ▶ A point estimator based on high-dimensional optimization
- ▶ Linear microlensing model (AIC) + a model w/o microlensing (BIC)

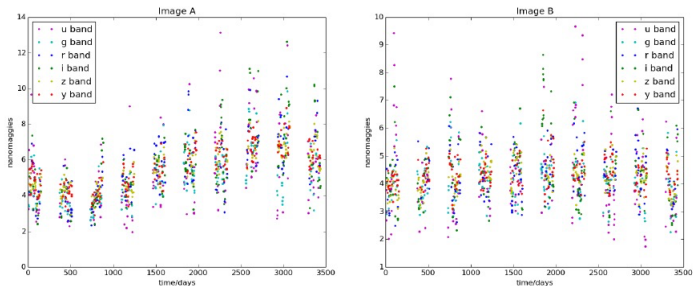
Kumar et al. (2014)

- ▶ A point estimator also based on high-dimensional optimization
- ▶ A spline with a Gaussian kernel to account for microlensing

It reveals that our model accounts for microlensing better than theirs.

# DISCUSSION: TIME DELAY CHALLENGE II

Doubly-lensed multi-filter light curves (Marshall et al., 2016+)



- ▶ Six bands,  $u, g, r, i, z, y$ , lead to vector time series.
- ▶ 5,000+ systems, 6 bands for each system, 150 obs. for each band.
- ▶ **Bluer** filters are more sensitive to microlensing than **redder** filters.
- ▶ All the information (except  $\Delta$ ) needed to calculate  $H_0$  will be given.
- ▶ Evaluation is based on two numbers,  $H_0$  estimate and its uncertainty.
- ▶ <http://timedelaychallenge.org> for more information!



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